

# 1 Analyse von Kontingenztafeln: Das loglineare Modell

## Aufgabe 12

Analyse des Datensatzes HairEyeColor

(a) Visualisierung der Datenstruktur

```
# ?HairEyeColor
class(HairEyeColor)

## [1] "table"

HairEyeColor

## , , Sex = Male
##
##      Eye
## Hair   Brown Blue Hazel Green
## Black   32   11   10    3
## Brown   53   50   25   15
## Red     10   10    7    7
## Blond    3   30    5    8
##
## , , Sex = Female
##
##      Eye
## Hair   Brown Blue Hazel Green
## Black   36    9    5    2
## Brown   66   34   29   14
## Red     16    7    7    7
## Blond    4   64    5    8

# Variablen
# 1 Hair
# 2 Eye
# 3 Sex

# Als "flache" Kontingenztafel
ftable(HairEyeColor)

##           Sex Male Female
## Hair Eye
## Black Brown      32      36
##       Blue      11       9
##       Hazel     10       5
##       Green      3       2
## Brown Brown     53      66
##       Blue     50      34
##       Hazel    25      29
##       Green    15      14
## Red   Brown     10      16
##       Blue     10       7
```

```
## Hazel 7 7
## Green 7 7
## Blond Brown 3 4
## Blue 30 64
## Hazel 5 5
## Green 8 8
```

```
# Randhäufigkeiten
apply(HairEyeColor,1,sum)
```

```
## Black Brown Red Blond
## 108 286 71 127
```

```
apply(HairEyeColor,2,sum)
```

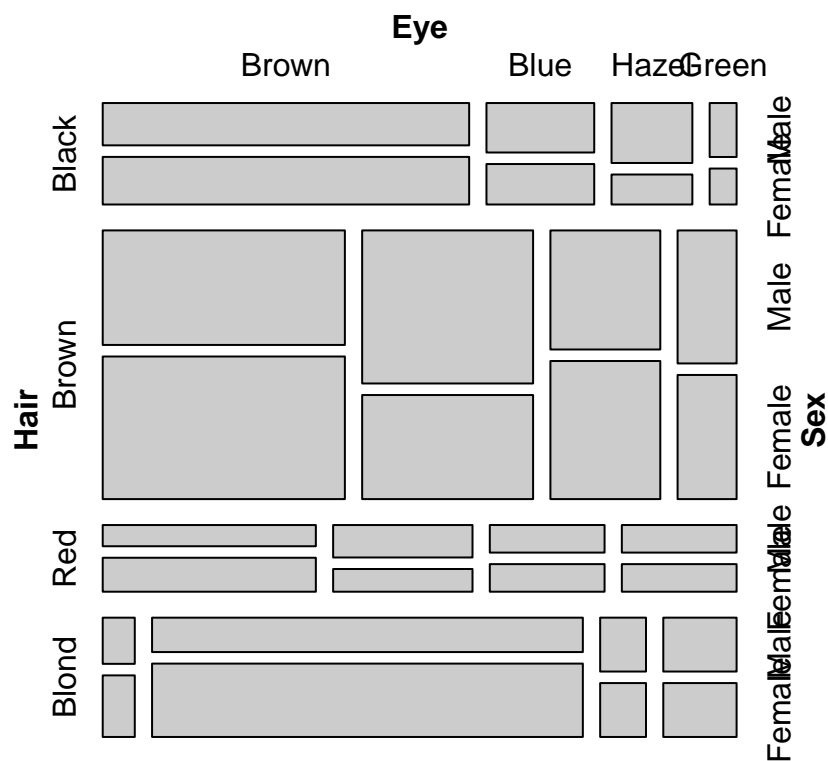
```
## Brown Blue Hazel Green
## 220 215 93 64
```

```
apply(HairEyeColor,3,sum)
```

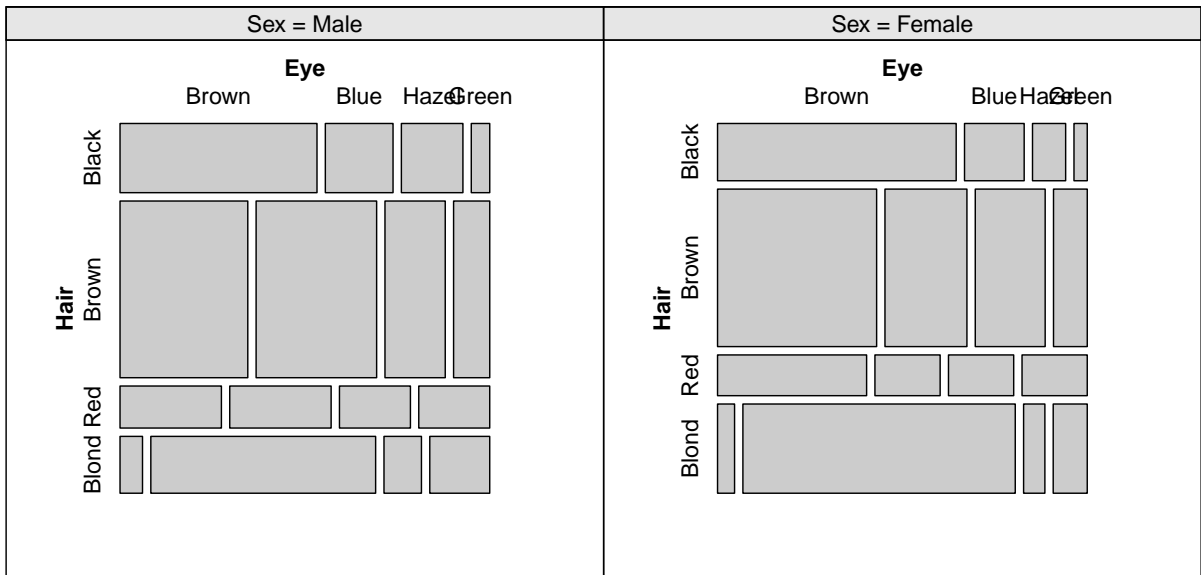
```
## Male Female
## 279 313
```

Paket vcd zur Visualisierung kategorialer Daten

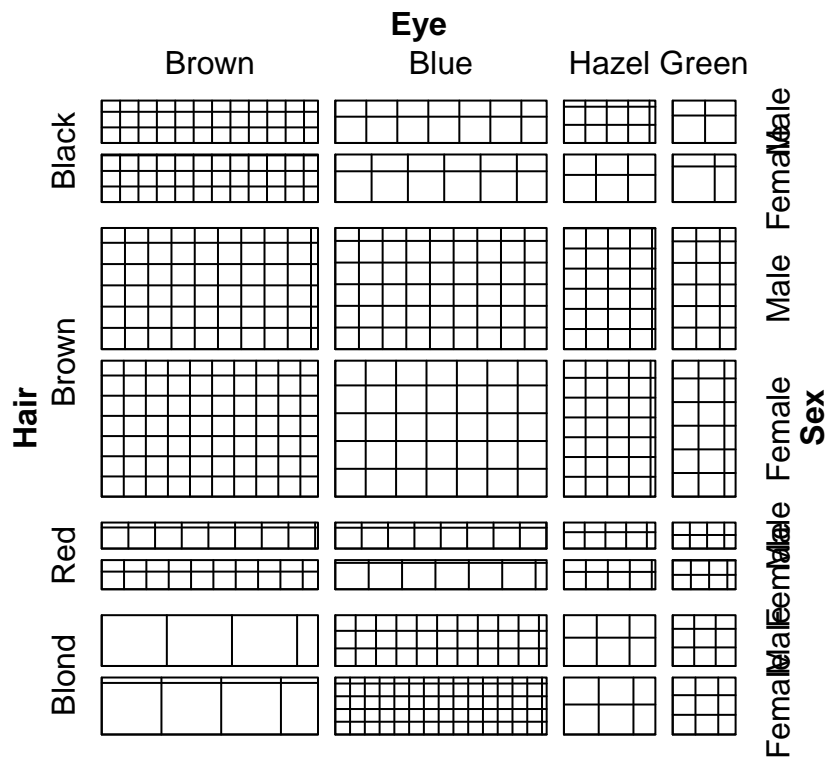
```
library(vcd)
# Darstellung der tatsächlichen Häufigkeiten
mosaic(HairEyeColor)
```



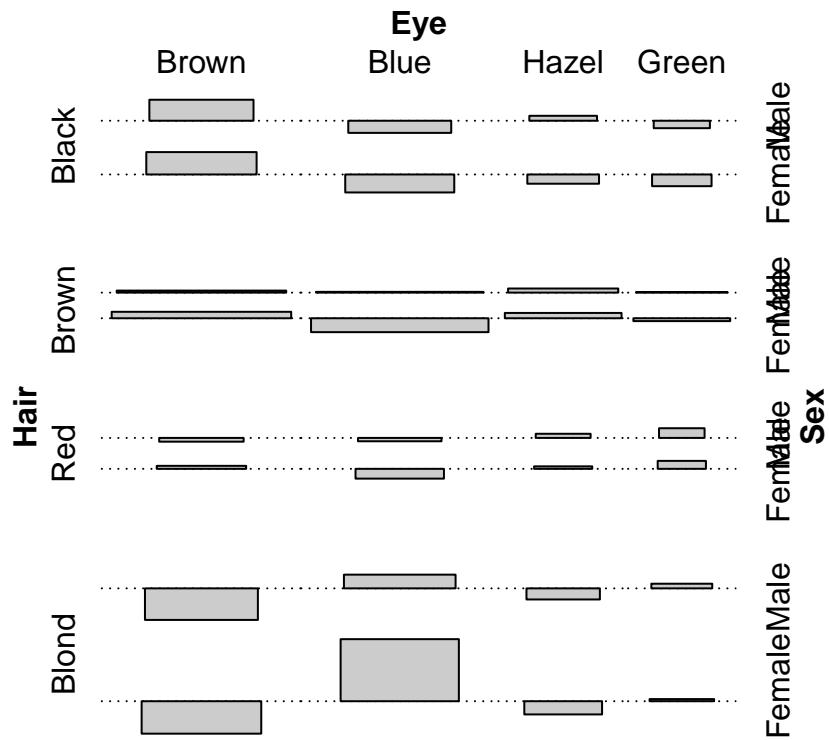
```
cotabplot(~ Hair + Eye | Sex, data=HairEyeColor)
```



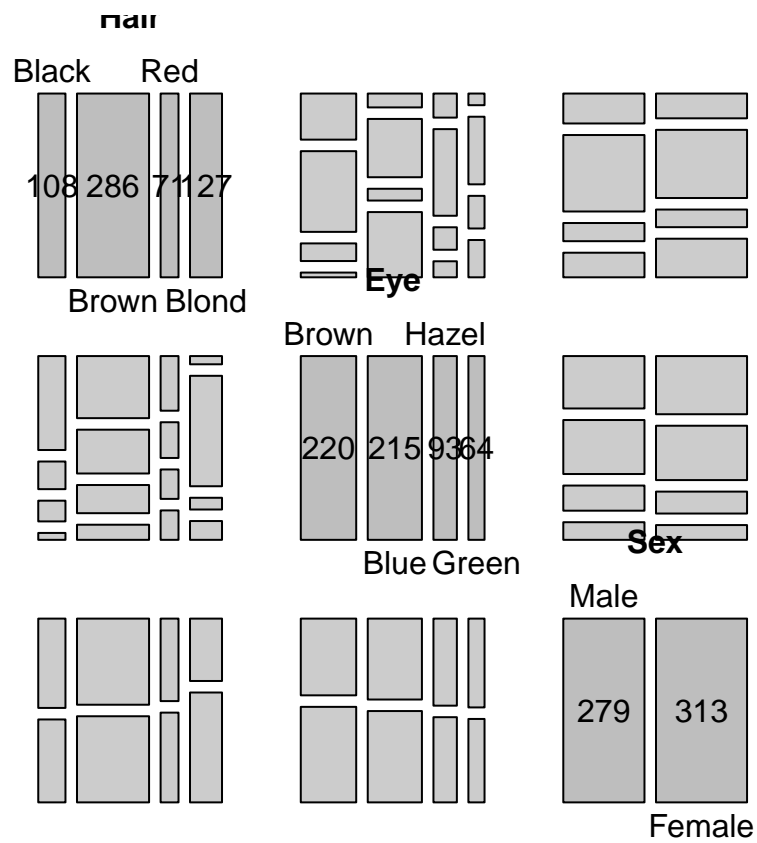
```
# Plots, die tatsächliche Häufigkeiten mit unter (totaler) Unabhängigkeit zu
# erwartenden Häufigkeiten vergleichen
sieve(HairEyeColor)
```



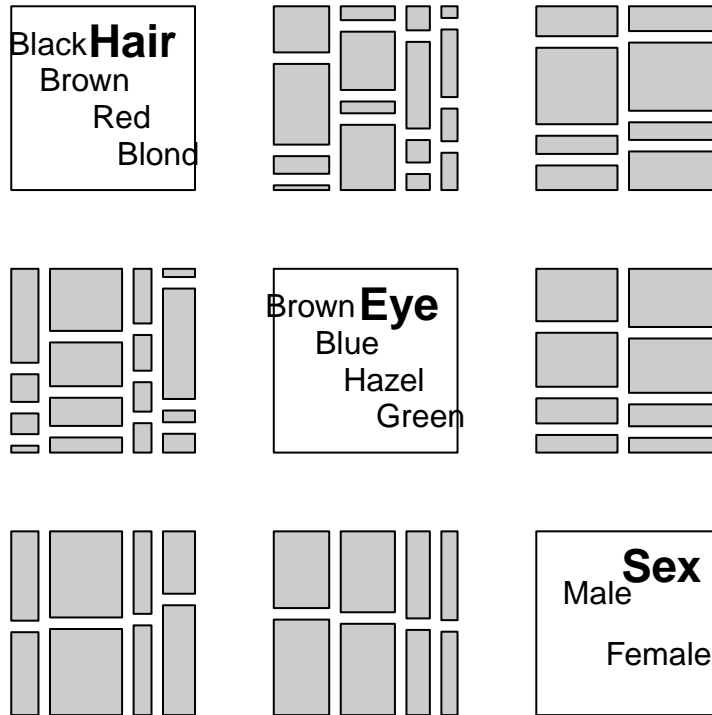
```
assoc(HairEyeColor)
```



```
# Kombination von paarweisen Plots
pairs(HairEyeColor)
```



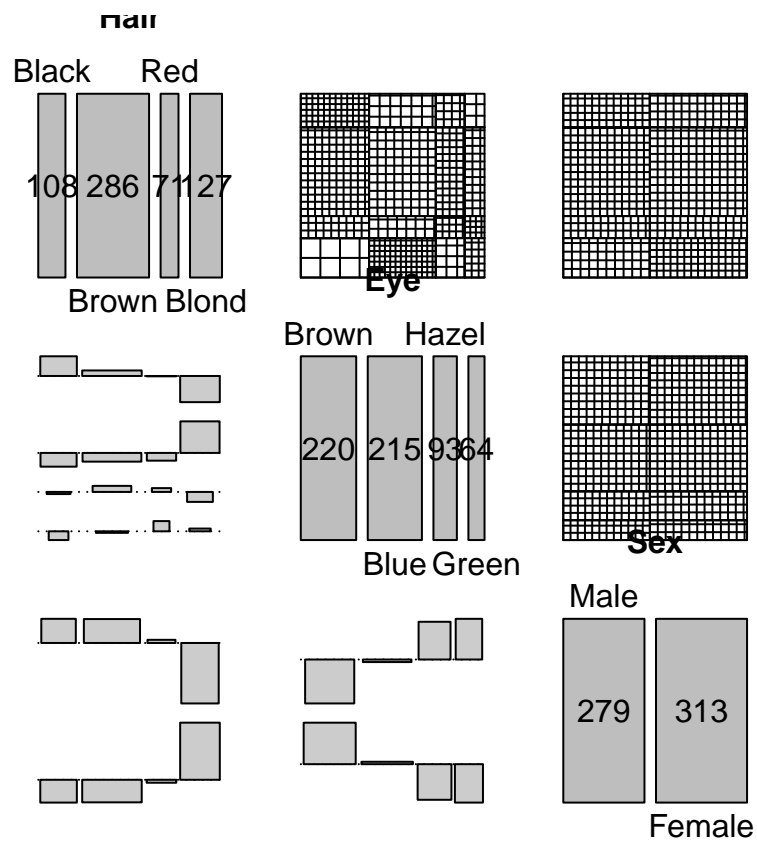
```
pairs(HairEyeColor, diag_panel=pairs_diagonal_text())
```



```

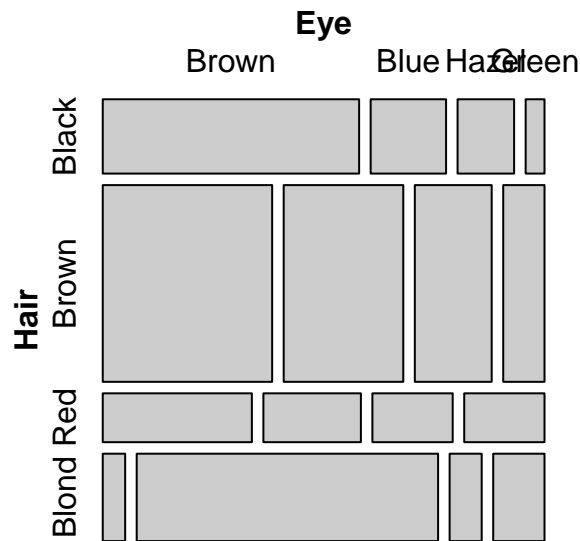
pairs(HairEyeColor, diag_panel=pairs_diagonal_mosaic(),
lower_panel=pairs_assoc(), upper_panel=pairs_sieve())

```



(b) Analyse ohne Variable *Sex*

```
# Aggregieren
HairEye <- HairEyeColor[, ,1] + HairEyeColor[, ,2]
mosaic(HairEye)
```



Loglineares Modell mit `loglin()`

```
# ?loglin
# Saturiertes modell
m0 <- loglin(HairEye, list(c(1,2)), fit=T, param=T)

## 2 iterations: deviation 0

m0

## $lrt
## [1] 0
##
## $pearson
## [1] 0
##
## $df
## [1] 0
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye"
##
##
## $fit
##      Eye
## Hair  Brown Blue Hazel Green
## Black   68  20  15   5
## Brown  119  84  54  29
## Red     26  17  14  14
## Blond   7   94  10  16
##
## $param
```

```

## $param$(Intercept)`
## [1] 3.189547
##
## $param$Hair
##      Black      Brown      Red      Blond
## -0.3063649  0.9520081 -0.3471908 -0.2984523
##
## $param$Eye
##      Brown      Blue      Hazel      Green
##  0.3611125  0.5112173 -0.2798778 -0.5924520
##
## $param$Hair.Eye
##      Eye
## Hair      Brown      Blue      Hazel      Green
## Black  0.9752132 -0.3986671  0.1047460 -0.6812921
## Brown  0.2764560 -0.2219556  0.1273068 -0.1818072
## Red    0.0546279 -0.5203601  0.0765790  0.3891532
## Blond -1.3062970  1.1409828 -0.3086318  0.4739461

# Haupteffekte
m1 <- loglin(HairEye, list(c(1),c(2)), fit=T, param=T)

## 2 iterations: deviation 0

m1

## $lrt
## [1] 146.4436
##
## $pearson
## [1] 138.2898
##
## $df
## [1] 9
##
## $margin
## $margin[[1]]
## [1] "Hair"
##
## $margin[[2]]
## [1] "Eye"
##
##
## $fit
##      Eye
## Hair      Brown      Blue      Hazel      Green
## Black  40.13514  39.22297  16.96622  11.675676
## Brown  106.28378  103.86824  44.92905  30.918919
## Red    26.38514  25.78547  11.15372  7.675676
## Blond  47.19595  46.12331  19.95101  13.729730
##
## $param
## $param$(Intercept)`
## [1] 3.341678
##
## $param$Hair
##      Black      Brown      Red      Blond
## -0.17911627  0.79474431 -0.59856762 -0.01706041
##
## $param$Eye
##      Brown      Blue      Hazel      Green
##  0.5296905  0.5067010 -0.3313375 -0.7050540

```

```

# Test auf Unabhängigkeit über Pearson-Statistik
m1$pearson>qchisq(0.95,9)

## [1] TRUE

# Test auf Unabhängigkeit über Devianz
m1$lrt>qchisq(0.95,9)

## [1] TRUE

```

→Jeweils, Vergleich der Teststatistik mit dem 95%-Quantil der Chi-Quadrat-Verteilung mit 9 Freiheitsgraden

→Unabhängigkeitshypothese wird in beiden Fällen abgelehnt

```

# Alternative: Berechnung der p-Werte:
1 - pchisq(m1$pearson,m1$df)

## [1] 0

1 - pchisq(m1$lrt,m1$df)

## [1] 0

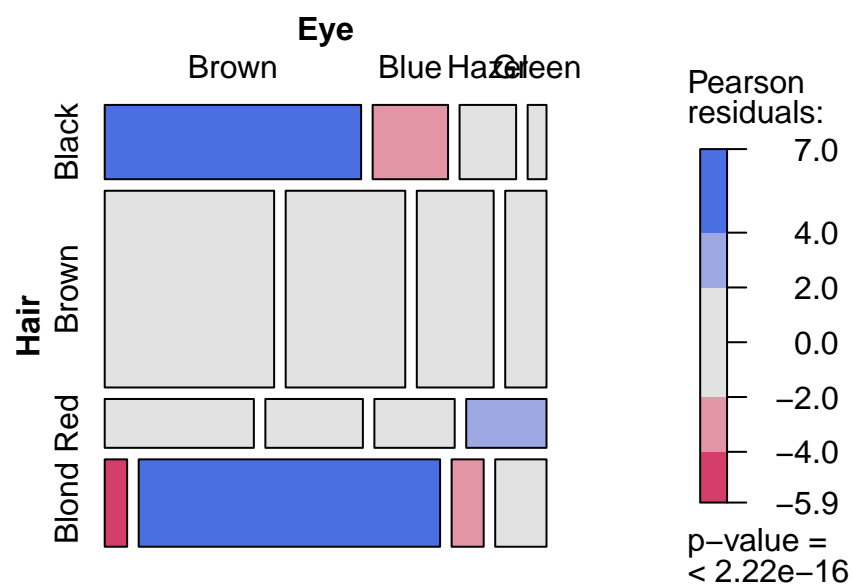
```

→Unabhängigkeitshypothese wird in beiden Fällen abgelehnt, da  $p$ -Werte  $\sim 0$ .

```

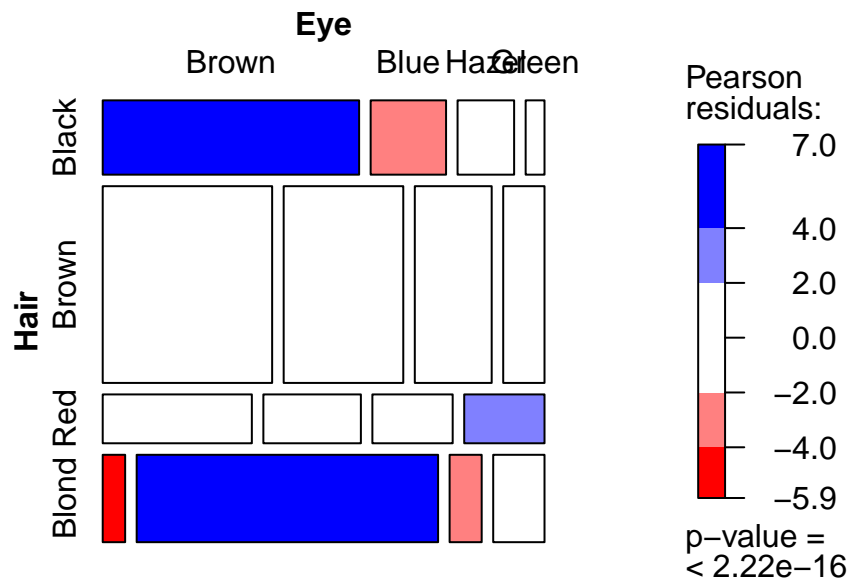
# Identifizierung der Zellen, die stark gegen Unabhängigkeitshypothese sprechen
mosaic(HairEye, shade=T) # shade=T äquivalent zu gp=shading_hcl, Defaultmethode

```

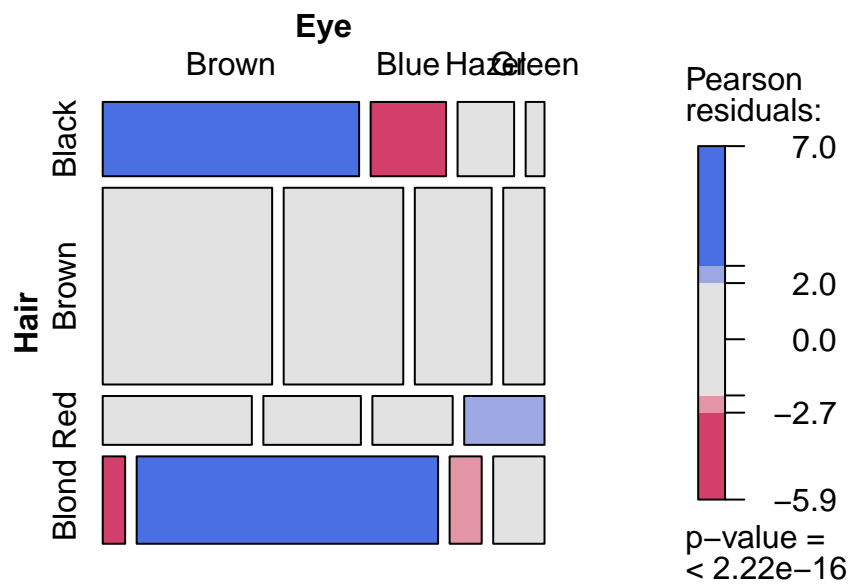




```
mosaic(HairEye, gp=shading_hsv)
```



```
mosaic(HairEye, gp=shading_max)
```



(c) Analyse der dreidimensionalen Tafel

```
# Saturiertes Modell
fm0 <- loglin(HairEyeColor, list(c(1, 2, 3)), fit=T)

## 2 iterations: deviation 0

fm0

## $lrt
## [1] 0
##
## $pearson
## [1] 0
##
## $df
## [1] 0
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye" "Sex"
##
##
## $fit
## , , Sex = Male
##
##      Eye
## Hair   Brown Blue Hazel Green
## Black   32   11   10    3
## Brown   53   50   25   15
## Red     10   10    7    7
## Blond    3   30    5    8
##
## , , Sex = Female
##
##      Eye
## Hair   Brown Blue Hazel Green
## Black   36    9    5    2
## Brown   66   34   29   14
## Red     16    7    7    7
## Blond    4   64    5    8

# Ohne 3-Faktor Interaktion
fm1 <- loglin(HairEyeColor, list(c(1, 2), c(1, 3), c(2, 3)), fit=T)

## 5 iterations: deviation 0.04093795

fm1

## $lrt
## [1] 6.761258
##
## $pearson
## [1] 6.868292
##
## $df
## [1] 9
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye"
```

```

##
## $margin[[2]]
## [1] "Hair" "Sex"
##
## $margin[[3]]
## [1] "Eye" "Sex"
##
##
## $fit
## , , Sex = Male
##
##      Eye
## Hair   Brown      Blue      Hazel      Green
## Black 32.789022 11.743604  8.443926  3.018481
## Brown 52.523050 45.931148 28.195250 16.348213
## Red   10.760485  8.819775  6.916602  7.502671
## Blond 1.927443 34.505472  3.444222  6.130636
##
## , , Sex = Female
##
##      Eye
## Hair   Brown      Blue      Hazel      Green
## Black 35.208836  8.258511  6.556304  1.981916
## Brown 66.477232 38.072151 25.804202 12.652225
## Red   15.239924  8.180608  7.083293  6.497414
## Blond 5.074007 59.488731  6.556201  9.868446

1 - pchisq(fm1$lrt, fm1$df) # H_0 wird beibehalten

## [1] 0.66196

# Bedingte Unabhängigkeiten
fm2a <- loglin(HairEyeColor, list(c(1, 2), c(1, 3)))

## 2 iterations: deviation 7.105427e-15

fm2a

## $lrt
## [1] 11.76372
##
## $pearson
## [1] 11.77059
##
## $df
## [1] 12
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye"
##
## $margin[[2]]
## [1] "Hair" "Sex"

1 - pchisq(fm2a$lrt - fm1$lrt, fm2a$df - fm1$df) # H_0 wird beibehalten

## [1] 0.1716167

fm2b <- loglin(HairEyeColor, list(c(1, 2), c(2, 3)))

```

```

## 2 iterations: deviation 1.421085e-14

fm2b

## $lrt
## [1] 18.32715
##
## $pearson
## [1] 18.0411
##
## $df
## [1] 12
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye"
##
## $margin[[2]]
## [1] "Eye" "Sex"

1 - pchisq(fm2b$lrt-fm1$lrt, fm2b$df-fm1$df) # H_0 wird abgelehnt

## [1] 0.009028299

fm2c <- loglin(HairEyeColor, list(c(1, 3), c(2, 3)))

## 2 iterations: deviation 2.842171e-14

fm2c

## $lrt
## [1] 156.6779
##
## $pearson
## [1] 147.944
##
## $df
## [1] 18
##
## $margin
## $margin[[1]]
## [1] "Hair" "Sex"
##
## $margin[[2]]
## [1] "Eye" "Sex"

1 - pchisq(fm2c$lrt-fm1$lrt, fm2c$df-fm1$df) # H_0 wird abgelehnt

## [1] 0

# 2 Variablen gemeinsam unabh. von der dritten
fm3a <- loglin(HairEyeColor, list(c(1, 2), c(3)))

## 2 iterations: deviation 5.684342e-14

fm3a

```

```

## $lrt
## [1] 19.85656
##
## $pearson
## [1] 19.56712
##
## $df
## [1] 15
##
## $margin
## $margin[[1]]
## [1] "Hair" "Eye"
##
## $margin[[2]]
## [1] "Sex"

1 - pchisq(fm3a$lrt-fm2a$lrt, fm3a$df-fm2a$df)           # H_0 wird abgelehnt

## [1] 0.04413148

fm3b <- loglin(HairEyeColor, list(c(1, 3), c(2)))

## 2 iterations: deviation 7.105427e-15

fm3b

## $lrt
## [1] 158.2073
##
## $pearson
## [1] 150.0845
##
## $df
## [1] 21
##
## $margin
## $margin[[1]]
## [1] "Hair" "Sex"
##
## $margin[[2]]
## [1] "Eye"

1 - pchisq(fm3b$lrt-fm2a$lrt, fm3b$df-fm2a$df)       # H_0 wird abgelehnt

## [1] 0

# -> Endmodell 1 2 | 1 3 also Hair Eye | Hair Sex

```

## Aufgabe 13

### Analyse der Unfalldaten

(a) Speichern der Daten als R-Objekt

```
accident <- data.frame(expand.grid(
  injury=factor(c("nonfatal", "fatal"), levels=c("nonfatal", "fatal")),
  ejected=factor(c("yes", "no"), levels=c("yes", "no")),
  safety=factor(c("belt", "none"), levels=c("belt", "none")),
  count=c(1105, 14, 411111, 483, 4624, 497, 157342, 1008))
accident

##      injury ejected safety  count
## 1 nonfatal     yes  belt   1105
## 2  fatal     yes  belt    14
## 3 nonfatal     no  belt 411111
## 4  fatal     no  belt   483
## 5 nonfatal     yes  none  4624
## 6  fatal     yes  none   497
## 7 nonfatal     no  none 157342
## 8  fatal     no  none  1008
```

(b) Loglineare Modell mit `loglm()`

```
library(MASS)
# ?loglm

# Saturiertes Modell
fm0 <- loglm(count ~ injury*ejected*safety, data=accident, param=T, fit=T)
# Modell ohne 3-Faktor Interaktion
fm1 <- update(fm0, ~ . - injury:ejected:safety)
# . steht für bisheriges Modell fm0: injury*ejected*safety,
# davon werden 3-Faktor-Interaktionen entfernt

# Geschätzte Parameter
fm0$param

## $(Intercept)
## [1] 7.78535
##
## $injury
## nonfatal fatal
## 2.299503 -2.299503
##
## $ejected
## yes no
## -1.711784 1.711784
##
## $safety
## belt none
## -0.5970266 0.5970266
##
## $injury.ejected
## ejected
## injury yes no
## nonfatal -0.6497609 0.6497609
## fatal 0.6497609 -0.6497609
##
```

```

## $injury.safety
##           safety
## injury      belt      none
## nonfatal  0.4792833 -0.4792833
## fatal     -0.4792833  0.4792833
##
## $ejected.safety
##           safety
## ejected      belt      none
##   yes -0.6532103  0.6532103
##   no  0.6532103 -0.6532103
##
## $injury.ejected.safety
## , , safety = belt
##
##           ejected
## injury      yes      no
## nonfatal  0.0552462 -0.0552462
## fatal     -0.0552462  0.0552462
##
## , , safety = none
##
##           ejected
## injury      yes      no
## nonfatal -0.0552462  0.0552462
## fatal     0.0552462 -0.0552462

```

fm1\$param

```

## $(Intercept)`
## [1] 7.831967
##
## $injury
## nonfatal      fatal
## 2.251692 -2.251692
##
## $ejected
##   yes      no
## -1.663278  1.663278
##
## $safety
##   belt      none
## -0.5489884  0.5489884
##
## $injury.ejected
##           ejected
## injury      yes      no
## nonfatal -0.6994479  0.6994479
## fatal     0.6994479 -0.6994479
##
## $injury.safety
##           safety
## injury      belt      none
## nonfatal  0.4293303 -0.4293303
## fatal     -0.4293303  0.4293303
##
## $ejected.safety
##           safety
## ejected      belt      none
##   yes -0.599909  0.599909
##   no  0.599909 -0.599909

```

```

# Gefittete Werte
fm0$fit

## , , safety = belt
##
##          ejected
## injury    yes    no
## nonfatal 1105 41111
## fatal     14    483
##
## , , safety = none
##
##          ejected
## injury    yes    no
## nonfatal 4624 157342
## fatal     497   1008

fm1$fit

## , , safety = belt
##
##          ejected
## injury    yes    no
## nonfatal 1098.13199 411117.868
## fatal     20.86801   476.132
##
## , , safety = none
##
##          ejected
## injury    yes    no
## nonfatal 4630.869 157335.131
## fatal     490.131  1014.869

```

Test über Devianz, ob Modell ohne 3-Faktor Interaktion passt

```

1 - pchisq(fm1$lrt, fm1$df)           # H_0 wird beibehalten

## [1] 0.09114565

```

Interpretation über bedingte odds ratios

```

OR <- function(X)
{
  return((X[1,1]*X[2,2])/(X[1,2]*X[2,1]))
}

# Verwende gefittete Werte:

# bedingt auf injury
fm1$fit[1,,]

##          safety
## ejected    belt    none
## yes      1098.132  4630.869
## no      411117.868 157335.131

fm1$fit[2,,]

```



```

##           safety
## ejected      belt      none
##      yes 20.86801 490.131
##      no 476.13203 1014.869

apply(fm1$fit,1,OR)

## nonfatal      fatal
## 0.090751 0.090751

# bedingt auf ejected
fm1$fit[,1,]

##           safety
## injury          belt      none
## nonfatal 1098.13199 4630.869
## fatal      20.86801 490.131

fm1$fit[,2,]

##           safety
## injury          belt      none
## nonfatal 411117.868 157335.131
## fatal      476.132 1014.869

apply(fm1$fit,2,OR)

##      yes      no
## 5.569589 5.569589

# bedingt auf safety
fm1$fit[, ,1]

##           ejected
## injury          yes      no
## nonfatal 1098.13199 411117.868
## fatal      20.86801 476.132

fm1$fit[, ,2]

##           ejected
## injury          yes      no
## nonfatal 4630.869 157335.131
## fatal      490.131 1014.869

apply(fm1$fit,3,OR)

##      belt      none
## 0.0609445 0.0609445

```

→Bedingte Odds Ratios zweier Variablen auf jeder Stufe der dritten gleich (in allen drei Fällen)!

```

# Verwende reale Werte:

# bedingt auf injury
fatal <- matrix(accident[accident$injury=="fatal",],4,nrow=2)

```

```

nonfatal <- matrix(accident[accident$injury=="nonfatal",][,4],nrow=2)
OR(fatal)

## [1] 0.05878751

OR(nonfatal)

## [1] 0.09145975

# bedingt auf ejected
yes <- matrix(accident[accident$ejected=="yes",][,4],nrow=2)
no <- matrix(accident[accident$ejected=="no",][,4],nrow=2)
OR(yes)

## [1] 8.483456

OR(no)

## [1] 5.452904

# bedingt auf safety
belt <- matrix(accident[accident$safety=="belt",][,4],nrow=2)
none <- matrix(accident[accident$safety=="none",][,4],nrow=2)
OR(belt)

## [1] 0.09273043

OR(none)

## [1] 0.05960426

```

→Beziehung von oben für reale Werte nur approximativ erfüllt!

(c) Analyse mit *glm()*

```

# Poisson-Modell
glm1 <- glm(count ~ injury*ejected*safety - injury:ejected:safety,
data=accident, family=poisson)
summary(glm1)

##
## Call:
## glm(formula = count ~ injury * ejected * safety - injury:ejected:safety,
##      family = poisson, data = accident)
##
## Deviance Residuals:
##      1      2      3      4      5      6      7      8
## 0.20704 -1.59987 -0.01071  0.31400 -0.10095  0.30951  0.01731 -0.21583
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      7.00137    0.02992   233.99 <2e-16 ***
## injuryfatal     -3.96315    0.06944  -57.07 <2e-16 ***
## ejectedno       5.92527    0.02996  197.76 <2e-16 ***
## safetynone      1.43913    0.03321   43.33 <2e-16 ***
## injuryfatal:ejectedno -2.79779    0.05526  -50.63 <2e-16 ***
## injuryfatal:safetynone  1.71732    0.05402   31.79 <2e-16 ***

```

```
## ejectedno:safetynone -2.39964 0.03334 -71.97 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
## Null deviance: 1.6249e+06 on 7 degrees of freedom
## Residual deviance: 2.8540e+00 on 1 degrees of freedom
## AIC: 93.853
##
## Number of Fisher Scoring iterations: 3
```

*# Vergleich gefittete Werte*

accident

```
##      injury ejected safety count
## 1 nonfatal      yes  belt  1105
## 2   fatal      yes  belt   14
## 3 nonfatal      no  belt 411111
## 4   fatal      no  belt  483
## 5 nonfatal      yes  none 4624
## 6   fatal      yes  none  497
## 7 nonfatal      no  none 157342
## 8   fatal      no  none  1008
```

glm1\$fit

```
##           1           2           3           4           5           6
## 1098.13193 20.86807 411117.86807 476.13193 4630.86807 490.13193
##           7           8
## 157335.13193 1014.86807
```

fm1\$fit

```
## , , safety = belt
##
##      ejected
## injury      yes      no
## nonfatal 1098.13199 411117.868
## fatal    20.86801  476.132
##
## , , safety = none
##
##      ejected
## injury      yes      no
## nonfatal 4630.869 157335.131
## fatal    490.131  1014.869
```

*# Vergleich Parameter*

glm1\$coef

```
##      (Intercept)      injuryfatal      ejectedno
##      7.001366      -3.963146      5.925269
##      safetynone      injuryfatal:ejectedno      injuryfatal:safetynone
##      1.439134      -2.797795      1.717321
##      ejectedno:safetynone
##      -2.399636
```

fm1\$param

```

## $(Intercept)`
## [1] 7.831967
##
## $injury
## nonfatal fatal
## 2.251692 -2.251692
##
## $ejected
## yes no
## -1.663278 1.663278
##
## $safety
## belt none
## -0.5489884 0.5489884
##
## $injury.ejected
## ejected
## injury yes no
## nonfatal -0.6994479 0.6994479
## fatal 0.6994479 -0.6994479
##
## $injury.safety
## safety
## injury belt none
## nonfatal 0.4293303 -0.4293303
## fatal -0.4293303 0.4293303
##
## $ejected.safety
## safety
## ejected belt none
## yes -0.599909 0.599909
## no 0.599909 -0.599909

```

→ Gefittete Werte sind identisch, Parameter unterscheiden sich aber, da loglm() Effektkodierung verwendet, glm() per Default jedoch Dummy-Kodierung mit Referenzkategorien safety=belt, ejected=yes, injury=nonfatal.

⇒ Poisson-Modell mit Effektkodierung:

```

# Modell mit Effektkodierung
glm2 <- glm(count ~ injury*ejected*safety - injury:ejected:safety,
            data=accident, family=poisson,
            contrast=list(safety=contr.sum(2), ejected=contr.sum(2), injury=contr.sum(2)))

# Vergleich Parameter
glm2$coef

## (Intercept) injury1 ejected1 safety1
## 7.8319672 2.2516913 -1.6632772 -0.5489882
## injury1:ejected1 injury1:safety1 ejected1:safety1
## -0.6994486 0.4293302 -0.5999089

fm1$param

## $(Intercept)`
## [1] 7.831967
##
## $injury
## nonfatal fatal
## 2.251692 -2.251692

```

```
##
## $ejected
##      yes      no
## -1.663278  1.663278
##
## $safety
##      belt      none
## -0.5489884  0.5489884
##
## $injury.ejected
##      ejected
## injury      yes      no
## nonfatal -0.6994479  0.6994479
## fatal    0.6994479 -0.6994479
##
## $injury.safety
##      safety
## injury      belt      none
## nonfatal  0.4293303 -0.4293303
## fatal    -0.4293303  0.4293303
##
## $ejected.safety
##      safety
## ejected      belt      none
##      yes -0.599909  0.599909
##      no  0.599909 -0.599909
```

# -> Modelle nun identisch